8/1/83

National Criminal Justice Reference Service

# ncjrs

This microfiche was produced from documents received for inclusion in the NCJRS data base. Since NCJRS cannot exercise control over the physical condition of the documents submitted, the individual frame quality will vary. The resolution chart on this frame may be used to evaluate the document quality.



MICROCOPY RESOLUTION TEST CHART NATIONAL BUREAU OF STANDARDS-1963-A

Microfilming procedures used to create this fiche comply with the standards set forth in 41CFR 101-11.504.

Points of view or opinions stated in this document are those of the author(s) and do not represent the official position or policies of the U.S. Department of Justice.

National Institute of Justice United States Department of Justice Washington, D.C. 20531

## **Rhodes** Associates

Criminal Justice Research Series

#### DOCUMENTATION FOR SAS-COMPATIBLE LOGIT, TOBIT AND MVPROBIT PROCEDURES

### by

### David A. Lombardero Frederick C. Nold

January 11, 1983 CJRS 5

#### U.S. Department of Justice National Institute of Justice

This document has been reproduced exactly as received from the person or organization originating it. Points of view or opinions stated in this document are those of the authors and do not necessarily represent the official position or policies of the National Institute of Justice.

Permission to reproduce this copyrighted material has been granted by

Public	Domain/	LEAA/NIJ
	/	

U.S. Department of Justice

to the National Criminal Justice Reference Service (NCJRS).

Further reproduction outside of the NCJRS system requires permission of the copyright owner.

This report and the computer programs it describes were developed under grant number 81-IJ-CX-0055 from the National Institute of Justice. Responsibility for defects and errors in the programs and documentation rests solely with the authors, not the NIJ. Rhodes Associates will be happy to receive comments and questions regarding the programs and documentation.

Rhodes Associates, 706 Cowper Street, 3rd floor, Palo Alto, California 94301 (415) 326 6246

### Introduction

#### Instructions for

Setting Up How LOGIT Invoking I Using Spec

#### Instructions for

Setting Up How TOBIT How to Inv How MVPROB How to Inv Missing Da Options for Options for

Appendix - Descr MVPRO

Examples and Illu

Example 1; Illustratio

Example 2: Illustratio

Example 3: Illustratio

### Table of Contents

4	Page
	1
r Using SAS LOGIT Program	2
PROC LOGIT Works LCGIT from SAS Sial Features	2 3 4 4
The second real se	14
o TOBIT or MVPROBIT Works Toke TOBIT from SAS BIT Works Toke MVPROBIT from SAS ta OF TOBIT and MVPROBIT OF TOBIT only Forecasting	14 15 16 16 16 17 17 21
BIT Models	27
ustrations	
A Simple LOGIT Job on 1 (Output of Example 1)	<sup>i)</sup> 7 8
A More Complex LOGIT Job on 2 (Output of Example 2)	11 12 0
PROC TOBIT and PROC MVPROPIT on 3 (Output of Example 3)	22 24

### INTRODUCTION

This report provides documentation for three SAS-compatible statistical procedures: LOGIT, TOBIT and MVPROBIT. We embedded these routines in SAS to facilitate use of these statistical techniques by the research community and to allow users to take advantage of SAS's data management and manipulation capabilities. While these routines are not as efficient as similar self-contained programs which we developed; <sup>1</sup> additional computation-related costs are offset by ease of use and flexibility.<sup>2</sup>

<sup>1</sup>Several other individuals have made contributions to the self-contained versions of these programs. Paul A. David and, in particular, John G. Lewis contributed to a selfcontained program for PROBIT analysis. Also, Takeshi Amemiya commented on our design of the Tobit routine and David Ross helped develop this documentation, as well as the program examples it contains.

 $^{2}$ We wish to encourage use of these programs. We offer them at the cost of reproduction and shipping. This cost is \$75 and should be sent to Rhodes Associates. We will prepare a tape containing the three programs and mail it along with a copy of this report. You should specify the type of tape you want and the system upon which the programs will be installed.

LOGIT is a program that produces maximum likelihood estimates of parameters of a Multinomial Logit model.<sup>3</sup> This statistical technique is appropriate for modelling a categorical dependent variable (the outcome of a trial) as a function of a linear combination of other variables.<sup>4</sup> Unlike MVPROBIT, which allows only two possible outcomes, LOGIT allows as many as eight possible outcomes.<sup>5</sup>

7

Ŧ

1

1

6

8

The LOGIT program is not a standard SAS procedure. In order to use PROC LOGIT, therefore, a special DD statement, which accesses the LOGIT program, must be included in the JCL: // JOB // EXEC SAS

SAS statements

<sup>3</sup>An elementary discussion of the use of Logit analyses can be found in Chapter 10 of R. Pindych and D. Rubinfeld, Econometric Models and Economic Forecasts, Second Edition, McGraw Hill, 1981. Also see Takeshi Amemiya, "Qualitative Response Models: A Survey," Journal of Economic Literature, 1981, Vol. 19, 1483-1536.

<sup>4</sup>This statistical model is also appropriate for problems in which cardinal or ordinal dependent variables are reduced to categories such as "greater than zero" and "less than or equal to zero."

<sup>5</sup>Assessing the effect a particular independent variable has on the system of probabilities estimated in Logit models when there are three or more outcomes requires care. In addition to the nonlinear nature of the relationship between the probabilities and variables, one must treat the problem as a system recognizing that the probabilities of the various outcomes add to che.

### Instructions for Using SAS LOGIT Program

-2-

#### Setting Up PROC LOGIT

#### //STEPLIB DD DSN=data.set.name,DISP=SHR

This DD statement must immediately follow the // EXEC SAS statement.

-3-

#### How LOGIT Works

LOGIT uses an iterative procedure to find maximum likelihood estimates of the parameters of a Multinomial Logit model. The procedure assigns the value zero to each of the parameters as initial values unless the user specifies other values with the INITIAL option (see below). A quasi Newton method, which uses only first derivatives, generates fairly accurate estimates of the parameters of the Multinomial Logit model which maximize the likelihood function. The program then switches to Newton's method to refine the solution. The estimated covariance matrix of the maximum likelihood estimators of the parameters is the inverse of the matrix of second derivatives of the negative of the log of the likelihood function evaluated at or near the final estimates.

One of the outcomes must be selected as the normalization outcome. The user may specify which outcome is selected by using the NORMOUTCOME parameter described below or allow the program to make the selection. The rule governing the selection by the program is given in the NORMOUTCOME section. The results, in terms of the probabilities estimated for the various outcomes for a particular set of independent variables, are the same, regardless of which outcome is selected for normalization.

<sup>6</sup>At Stanford's computing center <u>data.set.name</u> is WYL.XA.U76.DLSASLIB . If you have transferred this program from tape to your facility, the operating system (OS) data set name that the program is stored under at your facility must be used.

Invoking LOGIT from SAS

3

6

**a** 

(f)

as follows: PROC LOGIT;

standard SAS default. Using Special Features

The syntax of the PROC LOGIT statement is:

PROC LOGIT options;

VAR variable list;

OUTCOME s;

Most applications do not require the special features that are available in LOGIT. In its simplest form, LOGIT is invoked

VAR INCOME AGE SEX;

OUTCOME BUYCAR;

The type of all independent variables must be numeric. If the VAR statement is omitted, LOGIT will use all numeric variables in the input data set (except the OUTCOME variable). This is the

The OUTCOME variable may be of any type; its value can be either a number or a string of up to eight characters.

LOGIT has two types of special features. The simplest features can be specified as options on the PROC LOGIT statement. The other more complicated features must be used in conjuction with a PARMCARDS4 statement.

1. The PROC LOGIT statement

The following options may appear in the PROC LOGIT statement: (a) NOINT

This option eliminates all intercepts from the model.

(b) EXACTCOV

This option will cause the covariance matrix to be computed using estimates of the independent variable coefficients generated at the final iteration rather than those from the preceeding iteration. This requires additional computing time and ordinarily is not necessary, since values in the last two iterations are quite similar, but may be desirable if the covariance matrix is to be used for further computations.

-5-

(c) NORMOUTCOME=s

Because the probabilities must add to 1, the probabilities of any m - 1 of the m outcomes determine the probability of the remaining outcome. Consequently, no coefficients are generated for one of the outcomes; this outcome is referred to as the normalization outcome. The NORMOUTCOME parameter allows the user to select the normalization outcome s. (If s is a character string, it must be enclosed in single quotes.) If this parameter is omitted, LOGIT selects the normalization outcome as follows: If there is an OUTCOME with a numeric value of zero, that becomes the normalization state. Otherwise, the alphabetically last (or numerically greatest) outcome is selected.

TOL=t (d)

> The value of t specifies the convergence criterion; the procedure will consider that it has found

(e) PRINT=n

This parameter specifies the name of the input SAS data set. If omitted, the last SAS data set created is the default. (This is the same as for standard SAS procedures.)

٢

 $\langle \hat{U} \rangle$ 

 $(\mathbf{I})$ 

sufficiently accurate estimates of the parameters of the model when the L<sub>2</sub> (euclidean) norm of the gradient is less than t. Ordinarily, the procedure selects a reasonable value of t that is based upon the number of parameters and the scale of the data. The value t may be stated in decimal form (e.g. .00001) or scientific notation of the form n.Dexp (e.g. 1.D-5).

This parameter controls the frequency of the printing of intermediate results and is primarily useful for diagnostic purposes. If n > 0, the initial estimate and each nth estimate thereafter are printed out. If n=0, or the PRINT statement is not used, only the results for the final guasi-Newton and each Newton iteration are printed out. The final estimates are printed regardless of the value of n.

#### (f) DATA=SASdsname

-6-

#### Example 1: A Simple LOGIT Job

This SAS job produces the output shown in Illustration 1:

//LOGITEST JOB // EXEC SAS //STEPLIB DD DSN=WYL.XA.U76.DLSASLIB,DISP=SHR //SAS.SYSIN DD \*

DATA; INPUT OWNHOME \$ INCOME NUMCHILD; CARDS; Y 25 2 20 Y 1 35 Ν 2 Y 55 1 Y 40 0 Ν 30 2 Ν 80 2 Y 3.5 4 75 Ν 2 15 Ν 1

PROC LOGIT NORMOUTCOME='N'; VAR INCOME NUMCHILD; OUT COME OWNHOME; TITLE PROBABILITY OF HOME OWNERSHIP AS A FUNCTION OF; TITLE2 INCOME AND NUMBER OF CHILDREN;

•

Þ

- 🖶

i. L

Ť;

-₹¥

# 1 /

(r)

٢

Ð

O

RESULTS OF LOGIT ANALYSIS -2\*LN(L) = 12.9745

NORMALIZATION OUTCOME IS N -- NO COEFFICIENTS ARE GENERATED FOR THIS VALUE OF OWNHOME

COEFFICIENTS FOR OWNHOME = Y

1 INCOME (VARIABLE NUMCHILD (VARIABLE 2

3 INTERCPT (VARIABLE

COVARIANCE OF COEFFICIENTS (INDEXED AS ABOVE)

2 1 1 1 1.12690-03 -1.56 2 | -1.56580-03 4.092

-7-

#### Illustration 1'

#### PROBABILITY OF HOME OWNERSHIP AS A FUNCTION OF INCOME AND NUMBER OF CHILDREN

#### MULTINOMIAL LOGIT ANALYSIS OF OWNHOME

10 OBSERVATIONS UTILIZED (OTHERS HAD MISSING OR INVALID DATA)

MEAN	VARIANCE	SCALE FACTOR
41.0000000	444.000000	72.9931504
1.7000000	1.0100000	3.4813790

|GRADIENT| = 1.5510-02 6.487379

	COEFFICIENTS	GRADIENT
)	-0.028688	-1.0539850-02
ŧ.	-0,140259	-3.0331250-03
)	1.397386	-1.096352D-02

### HESSIAN (SCALED INVERSE COVARIANCE MATRIX) EVALUATED AT [GRADIENT] = 1.551D-02

[GRADIENT] = 5.231D-05

COEFFICIENTS	GRADIENT
-0.028299	3.5402360-05
-0.144108	1.951006D-05
1.393302	3.3199130-05

		ESTIMATED COEF	EST. STD. ERR.	EST. T-VALUE
ST.	1)	-0.02830	0.03357	-0.84299
1)	2)	-0.14411	0.63971	-0.22527
	3)	1.39330	1478653	0.77989

and a second second

3
58D-03 -4.2072D-02
220-01 -6.35370-01

3 | -4.2072D-02 -6.3537D-01 3.1917D+00

#### 2. The PARMCARDS4 statement

LOGIT recognizes certain statements that look like SAS statements but which SAS itself cannot process. These statements must be preceded by the SAS statement PARMCARDS4. (This is similar to the use of PROC BMDP.) The special LOGIT statements must appear in the order below (but all need not be present). They must be followed by a separate SAS statement consisting of four semicolons (;;;;).

-9-

(a) EXCLUDE s  $X_p \dots X_q$ ;

where s is an outcome other than the normalization outcome and  $X_p \ldots X_q$  is a list of one or more of the independent variables (the order of the X's is immaterial). The effect of this statement is to eliminate the terms corresponding to  $X_p \ldots X_q$  from the likelihood function for outcome s and is equivalent to setting the coefficients of these variables to zero. This feature may be utilized when  $X_p \ldots X_q$  are known not to affect the probability that the outcome is s. A different EXCLUDE statement is required for each outcome for which variable exclusion is to occur, but the order is not significant.

(b) INITIAL; or INITIAL (format);

This statement is utilized when the user wants to specify starting values for the coefficients. The string (format), if present, should be a valid FORTRAN format of at most 64 characters containing no semicolons.

•

3

3

3

3

ł

fre e

1

If the format is not specified, a default format of (5D16.8) (scientific notation, 16 characters per coefficient) is used. The coefficients must be in card format starting right after the START statement. The coefficients must appear in the order  ${}^{\beta}11' {}^{\beta}12 \cdots {}^{\beta}1n' {}^{\beta}10' {}^{\beta}21 \cdots {}^{\beta}2n' {}^{\beta}20' \cdots {}^{\beta}m-1, 1 \cdots {}^{\beta}m-1, 0$  where the  $\beta$ 's denote the coefficients, the first subscript the outcome and the second subscript the independent variable. The zeroth independent variable is the intercept. Any coefficients that are not utilized, either because they are covered by an EXCLUDE statement or because NOINT was specified, must be omitted.

(c) OUTPUT; or OUTPUT (format);

This statement, which is analogous to the START statement, is utilized to save the estimated coefficients and the covariance matrix for future use. The coefficients followed by successive rows of the covariance matrix, are written to FORTRAN file 20 (FT20F001). The user must include a JCL DD statement to define the attributes of file FT20F001, and (format) must be suitable for writing to that file. Remember to include a separate ;;;; statement at the end of all the special PARMCARDS4 statements.

-10-

### Example 2: A More Complex LOGIT Job

This SAS job produces the output in Illustration 2. The data

read by the job, contained in WYL.XA.U76.EDUC, is listed for

-11-

reference:

//LOGITST2 JOB // EXEC SAS //STEPLIB DD DSN=WYL.XA.U76.DLSASLIB,DISP=SHR //SURDATA3 DD DSN=WYL.XA.U76.EDUC,DISP=SHR

DATA SURVEY; INFILE SURDATA3; INPUT COLDEG \$ SATSCORE INCOME HSGRADES;

PROC LOGIT NORMOUTCOME='NONE' NOINT DATA=SURVEY; VAR SATSCORE INCOME HSGRADES; OUTCOME COLDEG;

PARMCARDS4; EXCLUDE GRAD HSGRADES;

WYL.XA.U76.EDUC

1721

1.	POSTGRAD	1210	22	3.4
2.	GRAD	980	35	3.6
3.	NONE	1020	25	3.2
4.	NONE	820	30	2.9
5.	GRAD	1100	44	3.1
6.	NONE	1140	30	2.8
7.	NONE	720	35	3.2
8.	POSTGRAD	1080	28	3.9
9.	NONE	1100	40	2.4
LO.	GRAD	1340	32	3.1
L1.	GRAD	1270	18	3.7
L2.	GRAD	990	24	3.8
L3.	POSTGRAD	1010	38	3.5
L4.	NONE	950	42	3.0
15.	GRAD	1280	38	3.7
L6.	NONE	920	16	2.6
17.	GRAD	1410	26	3.1
18.	POSTGRAD	1420	22	3.8
19.	GRAD	1150	48	3.9
20.	POSTGRAD	970	36	4.0

Note: In this program example, DATA=SURVEY is included as an option in the PROC LOGIT statement to demonstrate the DATA= option. However, since SURVEY is the most recently created SAS data set, it would have been used by default.

LOGI	T CONTROI	L CARD:	EXCL
	20 OF	20 OBS	SERVAT
NUMB OUTC	ER OF POS Omes (VA)	SSIBLE OUT LUES OF CO	COMES DLDEG
1		SATSCORE	1094
2		INCOME	31
3		HSGRADES	3
CONV	ERGENCE (	RITERIA S	ATISF:
FINA	L ITERATI	ION (NO.	6);
1	(GRAD	,SATSCOR	(E)
2	(GRAD	INCOME	)
3	( POSTGRA	D,SATSCOR	E)
4	( POSTERA	D, INCOME	)
5	( POSTGRA	D,HSGRADE	(S)
HESS	IAN (SCAL	ED INVERS	E COVA
NEWT	ON STEP N	10. 1;	FUNCT
1	(GRAD	,SATSCOR	E)
2	(GRAD	, INCOME	)

A.

3

Ľ

32

X

g .

6

8

3 (POSTGRAD, SATSCORE) 4 (POSTGRAD, INCOME )

5 (POSTGRAD, HSGRADES)

### Illustration 2

STATISTICAL ANALYSIS SYSTEM

MULTINOMIAL LOGIT ANALYSIS OF COLDEG

CLUDE GRAD HSGRADES;

VATIONS UTILIZED (OTHERS HAD MISSING CR INVALID DATA)

MES IS 3

EG ) ARE: GRAD POSTGRAD NONE

MEAN	VARIANCE	SCALE FACTUR
94.0000000	33444.0000000	633.5045383
51.4500000	74.4475000	29.8892957
3.3350000	0.2042750	1.5656628

ISFIED

);	FUNCTION VALUE	IS 19.3099	9   GRADIENT	= 1.2780-02
	COEFFICIENTS	GRADIENT		
	0.001609	1.7979890-0	3	
	-0.044093	-6.970026D-0	3	
	-0.002062	6.0381890-0	3	
	-0.131802	6.786470D-0	3	
	1.802093	5.3776580-0	3	

COVARIANCE MATRIX) EVALUATED AT |GRADIENT| = 1.278D-02

TION VALUE IS	19.30988	GRADIENT!	= 7.8170-05
COEFFICIENTS	GRADIENT		
0.001593	-3.0692040-05		
-0.043585	-1.5772950-05		
-0.002090	4.2488830-05		
-0.132215	3.5545690-05		
1.813591	4.3024180-05		

13

STATISTICAL ANALYSIS SYSTEM

RESULTS OF LOGIT ANALYSIS -2\*LN(L) = 38.6198

COEFFICIENTS FOR COLDEG= GRADESTIMATED COEFEST. STD. ERR.1SATSCORE (VARIABLE 1)0.001590.001612INCOME (VARIABLE 2)-0.043580.05145-0.84713	NOR	MALIZATION	OUTCOME	IS I	IONE		NO CI	DEFFICIENTS	ARE	GENERATED	FOR	THIS	VALUE	OF	COLDEG
ESTIMATED COEF     EST. STD. ERR.     EST. T-VALUE       1     SATSCORE (VARIABLE 1)     0.00159     0.00161     0.98827       2     INCOME (VARIABLE 2)     -0.04358     0.05145     -0.84713		COEFFICIE	NTS FOR	COLI	DEG	= GRAD	•								
1     SATSCORE (VARIABLE 1)     0.00159     0.00161     0.98827       2     INCOME (VARIABLE 2)     -0.04358     0.05145     -0.84713						EST	IMAT	ED COEF	EST	. STD. ERR	•	EST	T-VA	LUE	
2 INCOME (VARIABLE 2) -0.04358 0.05145 -0.84713	1	SATSCORE	(VARIAB	BLE	1)	C	.001	59	0	.00161		0	.98827		
	2	INCOME	(VARIAB	LE	2)	-0	. 043	58	0	.05145		-0	84713		
								÷.							
COEFFICIENTS FOR COLDEG = POSTGRAD	×	COEFFICIE	NTS FOR	COLI	DEG	= POST	GRAD								
ESTIMATED COEF EST. STD. ERR. EST. T-VALUE		1997 - A.				EST	IMAT	ED COEF	EST	. STD. ERR	•	EST	T-VA	LUE	
3 SATSCORE (VARIABLE 1) -0.00209 0.00321 -0.65130	3	SATSCORE	(VARIAB	LE	1)	-0	.002	09	0	.00321		-0	.65130		
4 INCOME (VARIABLE 2) -0.13222 0.07817 -1.69140	4	INCOME	(VARIAB	BLE a	2)	-0	.132	22	0	.07817		-1	.69140		
5 HSGRADES (VARIABLE 3) 1.81359 1.22260 1.48339	5	HSGRADES	(VARIAE	BLE	3)	1	.813	59	1	.22260		1	483'39		

COVARIANCE OF COEFFICIENTS (INDEXED AS ABOVE)

		1	2	3	<u>(</u> )	- 5	
			************				-
1	Ĩ	2.59830-06	-7.89590-05	1.83860-06	-5.20750-05	-1.54320-05	¢
,	1	-7 89590-05	2 64710-03	-5 36430-05	1.65150-03	3 92820-04	
•	1	-7.07270-02	2104/10-03	-2130-30 03	1103120 03	3.72025 0	
3	1	1.83860-06	-5.36430-05	1.03010-05	1.6063D-05	-3.21990-03	
4	1	-5.20750-05	1.65150-03	1.60630-05	6.1104D-03	-5.75160-02	
,	i						
5	I.	-1.54320-05	3,92820-04	-3.21990-03	-5.7516D-02	1.49470+00	

explanatory variables. Setting up TOBIT or MVPROBIT

1

Ł

\*

ŧ

**t** 

3

٢

8 × 1

1 :

¢ .

64

C

(1 ·

1

<sup>7</sup>For a discussion of Tobit models, see Takeshi Amemiya, "Regression Analysis when the Dependent Variable Is Truncated Normal," Econometrica, 1973, Vol. 41, 997-1016, and for Probit models, see Takeshi Amemiya, "Qualitative Response Models: A Survey," Journal of Economic Literature, 1981, Vol. 19, 1483-1536.

### Instruction for Using TOBIT and MVPROBIT Programs

Both Tobit and Probit<sup>7</sup> models rest on an underlying random variable which is normally distributed. The TOBIT program described below is appropriate for a normally distributed dependent variable whose mean is a linear combination of independent variables but where the observations have been truncated from below at zero. In other words, the dependent variable is not permitted to assume negative values; if its value would be negative, the value zero is observed instead.

The MVPROBIT program is appropriate for a dependent variable which can take on only 2 values, such as 0 and 1. The probability that the dependent variable takes on a particular value is assumed to be a standard normal cumulative distribution function (mean 0 and variance 1) whose argument is a linear combination of the

The TOBIT and MVPROBIT programs calculate maximum likelihood estimators of the parameters for Tobit and Probit analysis along with the corresponding asymptotic covariance matrix.

Neither TOBIT or MVPROBIT is a standard SAS procedure. In order to use these procedures, therefore, a special DD statement, which accesses the TOBIT (MVPROBIT) program, must be included in the JCL:

// JOB

// EXEC SAS

//STEPLIB DD DSN=data.set.name,DISP=SHR

SAS statements

This DD statement must immediately follow the // EXEC SAS statement.

#### How TOBIT Works

TOBIT uses an iterative procedure to find maximum likelihood estimates of the parameters of the model. The user may select the initial estimates by using the INITIAL option (see below). If the initial estimates are not specified, the program starts with the intercept set to the mean of Y, the outcome variable, the variance equal to the sample variance of Y, and the remaining parameters set to zero. The program uses a guasi-Newton method to obtain a fairly accurate estimate of the values of the parameters of the model which maximize the likelihood function.<sup>9</sup> The program then switches to Newton's method to refine the solution. The estimated covariance matrix is the inverse of the matrix of second derivatives of the negative of the log of the likelihood function evaluated at or near the final estimates.

<sup>8</sup>At Stanford's computing center data.set.name is WYL.XA.U76.DLSASLIB . If you have transferred this program from tape to your facility, the OS data set name that the program is stored under at your facility must be used.

<sup>9</sup>Internally, TOBIT scales the input data and coefficients in order to make the various parameters associated with each variable approximately equal in weight for purposes of assessing convergence. This improves numerical stability and makes convergence more rapid.

How to Invoke TOBIT from SAS The syntax of the PROC TOBIT statement is: PROC TOBIT options; VAR variable list; OUTCOME y; The type of the dependent variable and the independent variables must be numeric.<sup>10</sup> TOBIT can be used for problems where the truncation point is some number other than zero by pre-processing the data. For example, if the truncation point is  $\tau$ , create a new variable  $W = Y - \tau$ , and use TOBIT with W as the dependent variable. Then add  $\tau$  to intercept. Problems where the truncation is from above-where we observe  $Y = \min (z, \theta) - c$ an be transformed by multiplication by -1 and subtraction of  $\theta$  from the estimated intercept. How MVPROBIT Works MVPROBIT uses the same computational methods as TOBIT. Unless the INITIAL option is used, the first iteration begins with all the parameters set to zero. How to Invoke MVPROBIT from SAS The syntax for the PROC MVPROBIT statement is: PROC MVPROBIT options; VAR variable list; FREQ r s; Variables named in variable list are the independent variables while r and s are the names of the variables giving the frequencies of <sup>10</sup>If the VAR statement is omitted, all numeric variables in the input data set, except the OUTCOME variable, are utilized. This is the standard SAS default.

ALL ALL DATE AND ALL AND ALL

-15-

3

3

\$

t

E.

£

2

.

61

11

 $(\mathbf{0})$ 

.

-16-

	-17-			
outcomes	for each observation. Each observation may represent		in a province of the second	computing
several	trials, so r gives the number of times that a particular			be desiral
outcome	was observed and s the number of times that the second		i print Angli angli ang	further co
outcome	occurred. If an observation represents one trial, then	. <sup>1</sup>	n, in the pro-	(c) INITIAL
either	r = 1 and $s = 0$ or vice versa. The type of all variables,	1		If this or
includin	g frequencies, must be numeric. <sup>11</sup>			the user t
				β (β, σ
Missing	Data	1.		of the und
TOBI	T ignores any observation for which the value of any			supplied
independ	ent variable is missing or for which the value of the			by the in:
dependen	t variable is missing or negative.	lj e		ββ
MVPR	OBIT ignores any observation for which any data is missing		North Bank	utilized
or eithe	r of the frequencies is negative.	<i>2</i>		unless the
<u>Options</u>	for TOBIT and MVPROBIT	A constant		(d) OUTPUT
The	following options or parameters may appear in the PROC TOBIT			This option
or PROC	MVPROBIT statement:			parameter
(a)	NOINT	₫`¥		or cards o
	This option will cause the intercept term, $\beta_0$ , to be			β1, β2,
	omitted.			first; th
(b)	EXACTCOV	<u>(</u>		(or colum
	This option will cause the covariance matrix to be			The forma
	computed at the final estimate of the parameters rather			per line
	than merely at a close estimate. This requires additional	¢,		unless th
. <del></del>		1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		the user
<sup>11</sup> If the	VAR statement is omitted, all numeric variables in the			JCL state
input da	ta set, except the FREQ variables, are utilized. This tandard SAS default.	6		statement
				must be a
		f s	1.	

ag time and ordinarily is not necessary, but may able if the covariance matrix is to be used for computations.

option is specified, MVPROBIT (TOBIT) will expect to supply initial estimates of the coefficients, for TOBIT, where  $\sigma$  is the standard deviation underlying normal distribution). Estimates are through the use of a PARMCARD statement followed initial values. The values are input in the order  $\beta_0$ ,  $(\beta_1, \dots, \beta_n, \beta_0, \sigma)$  for TOBIT). The format is 5D15.0 (floating point, one every 15 columns), the INFORMAT parameter is specified.

tion causes MVPROBIT (TOBIT) to write the final er estimates and covariance matrix to disk, tape s on FORTRAN file 20 (FT20F001). The parameters  $\dots, \beta_n, \beta_0$  (and  $\sigma$  if TOBIT is used) are written then the covariance matrix is written one row umn, since the matrix is symmetric) at a time. mat used is (1P5D15.7), which provides 5 values e in scientific notation with 7 decimal places, the OUTFORMAT parameter is specified. Note that r must define the FT20F001 file using JCL. The tement should appear just after the //STEPLIB DD nt. If the default format is used, record length at least 75 characters. (e) TOL=t

The value of t specifies the convergence criterion; the procedure will consider that it has found sufficiently accurate estimates of the parameters of the model when the  $L_2$  (euclidean) norm of the gradient is less than t. Ordinarily, the procedure selects a reasonable value of t that is based upon the number of parameters and the scale of the data. The value t may be stated in decimal form (e.g. .00001) or scientific notation of the form n.Dexp (e.g. 1.D-5).

-19-

(f) PRINT=n

This parameter controls the frequency of the printing of intermediate results and is primarily useful for diagnostic purposes. If n > 0, the initial estimate and each nth estimate thereafter are printed out along with the final estimates. If n = 0, or PRINT is not used, only the results for the final quasi-Newton and each Newton iteration are printed out.

(g) INFORMAT=s

This parameter defines the input format when the INITIAL option is selected. The parameter s must be a string of at most 64 characters that comprise a valid FORTRAN format for reading cards, <u>e.g.</u>, (8F10.3).

(h) OUTFORMAT=s

This parameter defines the output format when the OUTPUT option is selected. The parameter must be a string of

3

3

**\$**.

3

3

P

44

**3** 

3

5.0

÷.

ą "

1

1 4

a i

1

**F** }

at most 64 characters that comprise a valid FORTRAN format for writing to file 20 (FT20F001). That file's characteristics must be defined by the user in a JCL statement. The OUTFORMAT parameter need not be specified if the default format is acceptable.

### DATA=SASdsname

(i)

This parameter specifies the name of the input data set. If omitted, the last data set created is the default. (This is the same as for other SAS procedures.) (j) STEPMAX=d

This parameter limits the total variation at each iteration in the parameter estimates for the model. Ordinarily, the value selected by the program  $(5+5\sqrt{n},$ where n is the number of parameters in the model), is suitable; however, in rare instances, it may be too large, resulting in computational problems such as underflows. If this occurs, the user should specify a smaller value of STEPMAX. This option should not be used except for ill-behaved problems since it will probably slow down convergence. The value d may be stated in decimal form (e.g. 10.0) or scientific notation of the form n.Dexp (e.g. 1.D+1).

-20-

#### Options for TOBIT only -- Forecasting

### (k) TESTDATA=SASdsname

This parameter specifies the name of the data set containing the test observations for which the expected values of the outcome variable and the variance of prediction and mean squared error of forecast of the estimate are to be computed. The variance of prediction is the estimate of variability for the estimate of the expected value of the outcome variable, and is appropriate for constructing confidence intervals for the expected value of the outcome variable. The mean squared error of forecast is the estimate of variability appropriate for normalizing comparisons of forecast values to corresponding realizations of the outcome variable. The TESTDATA data set must contain all of the independent variables in the VAR list, but not necessarily in the same order, and may contain the outcome variable and other variables which are not used in the analysis. If no TESTDATA data set is specified, no predicted values are computed. In the output, the value of -1.0 for the observed value of the dependent variable signifies that it was missing.

-21-

(1) PRINTTESTDATA

This option causes the values of the independent variables to be printed along with the forecasts.

(m)	TESTOUTF
	This par
	to FORTR
	number,
	and mean
	vation n
	data set
	The form
	independ
	immediat
	is (6X,
	for the
(n)	NOPRINTF
ð,	This opt
	suppress
Example	3: PROC
	//TOBTSI // EXEC //STEPLI //SAS.SY
	DATA CRI INPUT CARDS; 22 20

47

39

28 62

31

34 35 41

26

21

32

64

52

38

22

28

17

23

18

9

.

3

3

t

3

\$

£.

2

L

.

1 i

53

'ILE=n or TOF=n

Tameter causes the forecast results to be written TAN file |n|. The items written are observation observed Y, predicted Y, variance of prediction, a squared error of forecast. Note that the obsernumber is the absolute position in the TESTDATA t; observations which are unusable <u>are</u> counted. That is (I6, 4D16.8). If n < 0, then the lent variables are written on a new record tely following each forecast record. The format 4D16.8). The user must include a DD statement output file.

ORECAST or NPF

tion causes the printed forecast output to be sed if a TESTOUTFILE is also specified.

MVPROBIT and PROC TOBIT

rl Job SAS IB DD DSN=WYL.XA.U76.DLSASLIB,DISP=SHR (SIN DD \* IME; AGE INCOME OWNHOME BURGED80 NOTBURGD AMOUNT; 0 0 Δ 1 0 0 60 0 0 Λ 0 1 1 0 1 0 Ω 1 0 0 1 Ω

	Illustration 3
	PROBABILITY OF BEING BURGLARIZED AS A FUNCTION OF
-23-	AGE, INCOME AND HOME OWNERSHIP
	MULTIVARIATE PROBIT ANALYSIS OF BURGED80 AGAINST AG
26 37 1 1 0 50 54 15 0 0 1 0	29 OF 29 OBSERVATIONS UTILIZED (OTHERS HAD MISSING OR INVALID DATA)
43 35 1 1 0 575	
	MEAN VARIANCE SCALE FACTOR
39 48 1 0 1 0	1 AGE 37.6206897 147.5457788 24.2936847
	3 OWNHOME 0.4137931 0.2425634 0.9850246
36 11 1 0 1 0	
	CONVERGENCE CRITERIA SATISFIED
39 64 1 0 1 0	
31 32 0 1 0 45	FINAL ITERATION (NO. 8); FUNCTION VALUE IS 9.461441   GRADIENT! =
34 22 0 0 1 0	COEFFICIENTS GRADIENT 1 (AGE ) -0.077640 -4.504860D-03
	2 (INCOME ) 0.035921 -1.596771D-03
	3 (OWNHOME ) 0.394170 -8.975401D-04 4 (INTERCET) 0.276199 -3.3933070-03
VAR AGE INCOME OWNHOME:	
FREQ BURGED80 NOTBURGD;	HESSIAN (SCALED INVERSE COVARIANCE MATRIX) EVALUATED AT [GRADIENT] = 5.930D-0
TITLE PROBABILITY OF BEING BURGLARIZED AS A FUNCTION OF; TITLE2 AGE, INCOME AND HOME OWNERSHIP:	
TITLE3 ILLUSTRATION OF PROC MVPROBIT;	The NEWTON STEP NO. 1; FUNCTION VALUE IS 9.461439  GRADIENT  = 2.461
DATA RESERVED:	COEFFICIENTS GRADIENT
INPUT AGE INCOME OWNHOME BURGED80 NOTBURGD AMOUNT;	1 (AGE ) -0.077615 1.933753D-06
CARDS; 26 23 0 1 0	3 (OWNHOME) 0.394174 3.097021D-07
	4 (INTERCPT) _ 0.276816 1.312641D-06
43 33 1 1 0 617	FINAL RESULTS OF PROBIT ANALYSIS OF BURGED80
$ \begin{array}{ccccccccccccccccccccccccccccccccc$	-2*LN(L) = 18.9229 NORM OF GRADIENT = 2.46CD-06
39 48 1 0 1 0	2 ESTIMATED COFF. EST. STD. FER. EST. T-VALUE
	1 AGE -0.07761 0.04459 -1.74082
36 11 1 0 1 0	2 INCOME 0.03589 0.02315 1.55063 3 OWNHOME 0.39417 0.84130 0.46353
	4 INTERCPT 0.27682 1.33057 0.20804
VAR AGE INCOME OWNHOME;	
OUTCOME AMOUNT;	COVARIANCE OF COEFFICIENTS
TITLE PREDICTED LOSS FROM BURGLARY AS A FUNCTION OF; TITLE2 AGE, INCOME AND HOME OWNERSHIP;	AGE INCOME OWNHOME INTERCPT
IITLES ILLUSTRATION OF PROC TOBIT;	+ <sup>™</sup>
	AGE   1.9878D-03 -2.7458D-04 -1.2616D-02 -4.8963D-02
	TNCOME -2 74580-04 5 35740-04 -3 31540-03 -4 33080-03
	OWNHOME   -1.2616D-02 -8.3156D-03 7.0778D-01 3.5726D-01
	INTERCPT   -4.89830-02 -6.33980-03 3.57250-01 1.77040+00

GE INCOME OWNHOME INTERCET

5.9300-03

50D-06



#### PREDICTED LOSS FROM BURGLARY AS A FUNCTION OF AGE, INCOME AND HOME OWNERSHIP ILLUSTRATION OF PROC TOBIT

#### FINAL RESULTS OF TOBIT ANALYSIS OF AMOUNT

NORM OF GRADIENT = 3.042D-05

ATED COEF	EST. STD. ERR.	EST. T-VALUE
86721	11.20484	-1.23761
35485	6.82276	0.93142
03330	229.86436	0.86587
34886	374.10233	-0.45268
31191	117.62554	2.87618

INCOME	OWNHOME	INTERCET	SIGMA
-1.7057D+01	-7.1362D+02	-2.5653D+03	-6.0667D+02
4.6550D+01	-5.9466D+02	-1.1680D+03	2.9847D+02
-5.9466D+02	5.2838D+04	1.3716D+04	3.5488D+03
-1.1680D+03	1.37160+04	1.39950+05	-7.2846D+03
2.98470+02	3.54880+03	-7.2846D+03	1.3836D+04

		<u>}</u>			
AMOUNT	VAR OF PRED	MSE FORECAST	AGE	INCOME	OHNHOME
61326	7462.6269	8293.4882	21.0000	18.0000	.0
67990	27074.330	31878.841	26.0000	37.0000	1.00000
14251	1716.1125	1796.4611	34.0000	13.0000	.0
59434	6985.5617	7573.8766	43.0000	33.0000	1.00000
93986	34616.255	45239.262	38.0000	73,0000	1.00000
13901	16293.275	17639.973	39.0000	48.0000	1.00030
48700	3509.2084	3750.6231	27,0000	13.0000	.0
72349	36493.177	43812.739	31.0000	60.0000	1.00000
21575	5321.3454	6228.4171	36.0000	11.0000	1.00000



. 🤅 🤊

ń :

( \*

0

\$

3

3

INITIAL option is used.

TOBIT

### -27-

#### APPENDIX

### Descriptions of the LOGIT, TOBIT and MVPROBIT Models

Let  $X_1, X_2, \ldots, X_n$  be independent variables and let R be a discrete dependent variable that can take on values  $s_1, s_2, \ldots, s_m$ (referred to as outcomes). Using independent observations on  $x_1, x_2, \ldots, x_n$  and R PROC LOGIT computes the maximum likelihood estimates for the Multinomial Logit model of the outcomes where the probability of outcomes depend on the coefficients,  $\beta_{\mbox{ij}},$  and

$$P\{R = s_{i}\} = \alpha_{i} / \sum_{j=1}^{m} \alpha_{j} ,$$

where  $\alpha_{i} = \exp(-\beta_{i0} - \sum_{j=1}^{n} \beta_{ij} X_{j}), \quad (j \neq m); \alpha_{m} = 1.$  The index m refers to the outcome selected for normalization. This selection can be made arbitrarily. Estimates of the standard errors and the asymptotic covariance matrix of the  $\{\beta_{ij}\}$  are also computed. LOGIT uses a two-stage computational algorithm to maximize the likelihood function. The procedure is iterative and starts with  $\beta_{ij} = 0$  for all i and j as the initial value unless the

Let  $X_1, X_2, \ldots, X_n$  be independent variables and Z be a variable that depends upon the X's. Furthermore, assume that wherever Z < 0, the observed value of the dependent variable is 0. Thus, the observed dependent variable is  $Y = \max (Z, 0)$ . Using the data for the X's and the Y's, TOBIT computes maximum likelihood estimators for the parameters,  $\beta_1$ ,  $\beta_2$ ,..., $\beta_n$ ,  $\beta_0$  and  $\sigma$ , when the density of Z is given by

-28-

$$Z \sim N(\beta_0 + \sum_{i=1}^{n} \beta_i x_i, \sigma^2)$$

where  $N(\mu,\sigma^2)$  is the univariate normal density with mean and variance  $\sigma^2$ . TOBIT also computes the asymptotic covariance matrix for these estimators. In addition, PROC TOBIT has an option which predicts the mean and variance of prediction and mean squared error of forecast of Y for a set of observation on the X's specified by the user.<sup>12</sup>

#### MVPROBIT

Let  $X_1, X_2, \ldots, X_n$  be independent variables and let S be a dichotomous variable that represents the observed outcome of a trial (for example, 1 = success, 0 = failure). Using independent observations on  $X_1, X_2, \ldots, X_n$  and S, PROC MVPROBIT computes maximum likelihood estimators for the binomial model of the outcome where the probability of the outcome depends on  $\beta_i$ 's and is given by

$$\Pr \{s = 1\} = F(\beta_0 + \sum_{i=1}^{n} \beta_i X_i),$$

<sup>12</sup>Formulas are available in Amemiya (1973) <u>op.cit</u>.and section 7.1 of <u>Continuous Univariate Distribution 1</u>, N.L. Johnson and S. Kotz, 1970, Houghton Mifflin Company, Boston. where F is the univariate normal cumulative distribution function with mean 0 and variance 1. (Note that the variance can be assumed to be 1 without loss of generality since the variance is incorporated as a scale factor into the  $\beta_i$ 's.)

e l

3

1990 1910

•

1

L

Ľ

1

M)

67

C.

nt

